
Pricing long-lived securities in dynamic endowment

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Abstract

We solve for asset prices in a general affine representative-agent economy with isoelastic recursive utility and rare events. Our novel solution method is exact in two special cases: no preference for early resolution of uncertainty and elasticity of intertemporal substitution equal to one. Our results clarify model properties governed by the elasticity of intertemporal substitution, by risk aversion, and by the preference for early resolution of uncertainty. Our results also highlight that a covariance-based factor structure arises as a very special case, rather than as a general property, of equilibrium models.

Introduction

The framework of representative-agent asset pricing, in which complete markets allows for the diversification of idiosyncratic risks, has for many years delivered benchmark models of the cross-section and time-series of stock prices and returns.¹ These models are at the same time simple and rich in the types of economic intuition they capture. In this paper, we focus on two nested sub-classes of the dynamic representative-agent framework with the goal of clarifying important implications for risk premia and asset prices. In the first part of the paper, we extend classic cross-sectional results of Merton (1973) and Breeden (1979) to a dynamic setting with recursive utility (Epstein and Zin, 1989) and rare events. This section assumes only isoelastic recursive utility and a Markov structure. Widening the class of models beyond the traditional diffusion framework of Merton (1973) has dramatic implications for the cross-section. The intertemporal capital asset pricing model (ICAPM) of Merton is a standard

justifications for the near-universal use of covariance-based factor models in finance. However, the ICAPM relies on conditional log-normality. Without the assumption of conditional log-normality, a factor structure may not hold

In this section, we also show that the dynamics of the wealth-consumption ratio are principally governed by the elasticity of intertemporal substitution (EIS) and the risk premia in the cross-section relative to a consumption-based model are governed by the preference for early resolution of uncertainty. However, relative to a wealth-based model, risk premia in the cross-section are governed by risk aversion. When risk aversion is equal to one, a rare-event wealth CAPM holds regardless of the EIS. In contrast, a rare-event consumption CAPM holds if there is no preference for early resolution of uncertainty, regardless of risk aversion. Some of these comparative statics have appeared elsewhere in the literature, but the advantage

of our framework is that we can show them in a more general setting.

In the second part of the paper, we derive approximate analytical solutions for the pricing of long-lived assets. Our solution method takes as its starting point the widely used method of Campbell and Shiller (1988) which involves a first-order approximation of the price-dividend ratio by a log-linear function. Previous studies use this method to compute the wealth-consumption ratio (which is necessary for computing other asset prices under recursive utility), and then to compute prices on other assets. While we use the method to compute the wealth-consumption ratio, we then, given the approximation, compute prices on other assets exactly. As a consequence, our method, unlike others, is exact both when the elasticity of intertemporal substitution is equal to one, and when utility is time-additive. The reason is that the approximation for the wealth-consumption ratio is exact when the EIS is equal to one, and a log-linear wealth-consumption ratio is not necessary for closed-form solutions for asset prices under time-additive utility

In an example, we extend the model of Wachter (2013) to a case of non-unitary EIS. We show that our technique is notably closer to the solution when the exact problem is solved numerically, than the standard approximation. Moreover, because our solution is closer to the true solution in a formal sense, it delivers insight into the economics of the problem.

The remainder of the paper proceeds as follows. Section 2 describes our general set-up and derives results for the cross-section. Section 3 describes the affine set-up with

analytical solutions. Section 4 quantitatively evaluates the solution method under an example economy.

2. General Model

2.1 Assumptions Let B_t be a unidimensional Brownian motion and BX_t an n -dimensional Brownian motion, such that B_t and BX_t are independent. For $j = 1, \dots, m$, let N_{jt} be independent Poisson processes. Consider functions $\mu_c : \mathbb{R}^n \rightarrow \mathbb{R}$, $\sigma_c : \mathbb{R}^n \rightarrow \mathbb{R}$, $\mu_X : \mathbb{R}^n \rightarrow \mathbb{R}^n$, $\sigma_X : \mathbb{R}^n \rightarrow \mathbb{R}^n$, and $\lambda : \mathbb{R}^n \rightarrow \mathbb{R}^m$. Assume the endowment follows the process

$$\frac{dC_t}{C_t} = \mu_c(X_t) dt + \sigma_c(X_t) dB_{ct} + \sum_{j=1}^m (e^{Z_{cj}t} - 1) dN_{jt}, \quad (1)$$

where X_t is a vector of state variables following the process

$$dX_t = \mu_X(X_t) dt + \sigma_X(X_t) dB_{Xt} + \sum_{j=1}^m Z_{Xj} dN_{jt}, \quad (2)$$

and where, for all $j = 1, \dots, m$ and t , Z_{cj} are scalar random variables and ZX_{jt} an $n \times 1$ vector of random variables. We assume the joint distribution of Z_{cj} and ZX_{jt} is time-invariant, and thus suppress the t subscript when not essential for clarity. The intensity for Poisson process N_{jt} is time-varying and given by $\lambda_j(x)$, the j -th element of $\lambda(x)$. We adopt the convention that BX_t , and therefore X_t , are column vectors and that σ_c is a row vector.

Consider an asset that pays cash flows determined by the outcome of a dividend process $dD_t = \mu_d(X_t) dt + \sigma_d(X_t) dB_{dt} + \sum_{j=1}^m e^{Z_{dj}t} dN_{jt}$. (3)

There may be many such assets, but because we will assume complete markets, it will suffice to consider each asset in isolation, and therefore it is not necessary to add a subscript to D_t . Similarly to the rare-event outcomes for consumption and for X_t , Z_{dj} is a random variable with time-invariant distribution for all j . In what follows, let $B_t = [B_t, BX_t]$. For a

generic function $h(C_t, D_t, X_t)$, define where E_{v_j} denotes expectations taken with respect to the joint distribution of $Z_{c_j}, Z_{d_j}, Z_{X_j}$, and conditional on information just prior to t .

Assume a representative agent with recursively-defined utility where ψ is the elasticity of intertemporal substitution and γ is risk aversion (Duffie and Epstein, 1992b). When $\gamma = 1/\psi$, the recursion in (5) is linear, and (4) reduces to the time-additive case. We are interested in two limiting cases of (5). When $\psi = 1$, (5) reduces to

2.2 General characterization of the solution

We first characterize the value function, the wealth-consumption ratio, and the state-price density in terms of the state variables. Here and in the remainder of the paper, we follow the convention that partial derivatives with respect to a vector are row vectors; for example, $\partial I/\partial X = [\partial I/\partial X_1, \dots, \partial I/\partial X_n]$. Proofs not given in the main text are contained in Appendix A. Proposition 1 (Value function). Suppose the representative agent's preference is defined by (4)–(6), where the consumption growth process follows (1) and the state variable process follows (2). In equilibrium, $J(C_t, X_t) = V_t$, where

$$J(C_t, X_t) = \frac{C_t^{1-\gamma} I(X_t)^{1-\gamma}}{1-\gamma}, \tag{8}$$

for $\gamma \neq 1$ and

$$J(C_t, X_t) = \log C_t + \log I(X_t)$$

for $\gamma = 1$, where $I(\cdot)$ satisfies the partial differential equation

$$\begin{aligned} & \frac{\beta}{1-\frac{1}{\psi}} \left[I(x)^{\frac{1}{\psi}-1} - 1 \right] + \mu_c(x) + \frac{1}{2} \text{tr} \left[\left(\frac{1}{I} \frac{\partial^2 I}{\partial x^2} - \frac{\gamma}{I^2} \left(\frac{\partial I}{\partial x} \right)^\top \left(\frac{\partial I}{\partial x} \right) \right) \sigma(x) \sigma(x)^\top \right] \\ & + \frac{1}{I} \frac{\partial I}{\partial x} \mu_X(x) - \frac{1}{2} \gamma \sigma_c(x)^2 + \frac{1}{1-\gamma} \sum_{j=1}^m \lambda_j(x) E_{v_j} \left[e^{(1-\gamma)Z_{c_j}} \left(\frac{I(x+Z_{X_j})}{I(x)} \right)^{1-\gamma} - 1 \right] = 0. \end{aligned} \tag{9}$$

for $\psi \neq 1$ and

$$\begin{aligned} & -\beta \log I + \mu_c(x) + \frac{1}{2} \text{tr} \left[\left(\frac{1}{I} \frac{\partial^2 I}{\partial x^2} - \frac{\gamma}{I^2} \left(\frac{\partial I}{\partial x} \right)^\top \left(\frac{\partial I}{\partial x} \right) \right) \sigma(x) \sigma(x)^\top \right] \\ & + \frac{1}{I} \frac{\partial I}{\partial x} \mu_X(x) - \frac{1}{2} \gamma \sigma_c(x)^2 + \frac{1}{1-\gamma} \sum_{j=1}^m \lambda_j(x) E_{v_j} \left[e^{(1-\gamma)Z_{c_j}} \left(\frac{I(x+Z_{X_j})}{I(x)} \right)^{1-\gamma} - 1 \right] = 0. \end{aligned} \tag{10}$$

We solve for asset prices in a general affine representative-agent economy with isoelastic recursive utility and rare events. Our novel solution method is exact in two special cases: no preference for early resolution of uncertainty and elasticity of intertemporal substitution equal to one. Our results clarify model properties governed by the elasticity of intertemporal substitution, by risk aversion, and by the preference for early resolution of uncertainty. Our results also highlight that a covariance-based factor structure arises as a very special case, rather than as a general property, of equilibrium models

for $\psi = 1$. Equation 10 is a special case of (9) as can be seen by taking the limit as $\psi \rightarrow 1$. Given the value function (8), we can now express the wealth-consumption ratio and the state price density in terms of $I(X_t)$. The consumption-wealth ratio will play an important role in our solution method for the affine case. Corollary 2 (Wealth-consumption ratio). Let W_t denote the wealth of the representative agent at time t . Then the wealth-to-consumption ratio $G_c(X_t) \equiv W_t/C_t$ is a function of X_t and is given by

$$G_c(X_t) = \begin{cases} \beta^{-1} I(X_t)^{1-\frac{1}{\psi}} & \psi \neq 1 \\ \beta^{-1} & \psi = 1. \end{cases} \tag{11}$$

Conclusions

The second and third terms on the right-hand side of (A.29) have zero expectation. Therefore

the first term in (A.29) must also have zero expectation, and it follows that the integrand of this term must equal zero. as can be seen by taking the limit as $\psi \rightarrow 1$. Given the value function (8), we can now express the wealth-consumption ratio and the state price density in terms of $I(X_t)$. The consumption-wealth ratio will play an important role in our solution method for the affine case. Corollary 2 (Wealth-consumption ratio). Let W_t denote the wealth of the representative agent at time t . Then the wealth-to-consumption ratio is useful

References

- Bansal, Ravi, and Amir Yaron, 2004, Risks for the long run: a potential resolution of asset pricing puzzles, *Journal of Finance* 59, 1481–1509.
- Barro, Robert J., and Jos'e F. Urs'ua, 2008, Macroeconomic crises since 1870, *Brookings Papers on Economic Activity* 39, 255–350.
- Benzoni, Luca, Pierre Collin-Dufresne, and Robert S. Goldstein, 2011, Explaining asset pricing puzzles associated with the 1987 market crash, *Journal of Financial Economics* 101, 552 – 573
- . Breeden, Douglas T., 1979, An intertemporal asset pricing model with stochastic consumption and investment opportunities, *Journal of Financial Economics* 7, 265–296.
- Campbell, John Y., 1993, Intertemporal asset pricing without consumption data, *American Economic Review* 83, 487–512.
- Campbell, John Y., Andrew W. Lo, and A. Craig MacKinlay, 1997, *The Econometrics of Financial Markets*. (Princeton University Press Princeton, NJ). Campbell, John Y., and Robert J. Shiller, 1988, The dividend-price ratio and expectations of future dividends and discount factors, *Review of Financial Studies* 1, 195–228.
- Chacko, George, and Luis Viceira, 2005, Dynamic consumption and portfolio choice with stochastic volatility in incomplete markets, *Review of Financial Studies* 18, 1369–1402.
- Cochrane, John H., 2001, *Asset Pricing*. (Princeton University Press Princeton, NJ).
- Drechsler, Itamar, and Amir Yaron, 2011, What's vol got to do with it, *Review of Financial Studies* 24, 1–45.
- Duffie, Darrell, 1996, *Dynamic Asset Pricing Theory*. (Princeton University Press Princeton, NJ).